

# The Association between Vitamin D and Hashimoto Thyroiditis: An Up-to-date Systematic Review and Meta-analysis

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## Abstract

The objective of the current study was to summarize the up-to-date studies to investigate the relationship between vitamin D and Hashimoto thyroiditis (HT). An online search of English and Chinese databases was performed. The studies concerned the investigation of the relationship between vitamin D and HT including meta-analysis, meanwhile the heterogeneities were revealed by subgroup analysis. Forty six related studies containing 15,336 participants (HT: 6,138 versus control: 9,198) were included. HT patients had lower levels of 25(OH)D<sub>3</sub> (standardised mean difference, -1.09; 95%CI: [-1.42, -0.75];  $P < 0.01$ ), and were more likely to be deficient in 25(OH)D<sub>3</sub> (OR, 2.77; 95%CI, [1.88, 3.91];  $P < 0.05$ ). Obvious heterogeneities in the results of meta-analysis were down to the difference of detection methods and criteria of vitamin D insufficiency among studies. Vitamin D deficiency was concluded to have a significant relation with HT.

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## INTRODUCTION

T and B lymphocyte mediated immune tolerance disorder led to abnormal auto-antibodies invasion in thyroid glands, which was the main mechanism of Hashimoto thyroiditis (HT). The increase of HT continued<sup>[1]</sup>, and gradually became the most common cause of thyroid hormone insufficiency. Besides, it was suggested a link to diseases including ischemic heart disease, osteoporosis, diabetes, cardiovascular disease and cancer<sup>[2-6]</sup>. Papillary thyroid carcinoma was the most common form of cancer associated with HT<sup>[7]</sup>. Even the health of children was threatened by HT with higher risk of dyslipidemia and cardiovascular disease<sup>[8]</sup>. Along with the progress of HT, the secretion of thyroid hormone was significantly insufficient, and patients had to replenish with Euthyroxin throughout their life, but still suffered from higher risk of HT related complications. Therefore, effective methods for HT prevention were considered to be urgent research.

Since Vitamin D receptors were shown to be present in the thyroid<sup>[9]</sup>, Vitamin D was considered as HT prevention in recent studies. Krysiak et al.<sup>[10]</sup> suggested that Vitamin D combined with atorvastatin could improve thyroid autoimmunity. In addition, supplementation with cholecalciferol for HT patients was indicated to twist the balance of CD4<sup>+</sup> T-cell subsets toward ameliorative composition<sup>[11]</sup>. Although the relationship between Vitamin D and HT seemed clear, it still needed further investigation<sup>[12]</sup>, for there existed study reported they had no evident correlation<sup>[13]</sup>. These contradictory conclusions maybe due to the small sample size of the local population, but it was very difficult to conduct a high-quality epidemiological study

with a large sample size. Therefore, to indicate whether vitamin D deficiency was really correlated to HT, which may affect clinical strategy whether we should use Vitamin D to prevent HT, evidence-based medical research tools such as systematic review and meta-analysis were necessary for a more reliable conclusion.

We conducted a systematic review and meta-analysis of the domestic and foreign studies that investigated the relationship between Vitamin D and HT. Compatible data were pooled into meta-analysis to provide rigorous evidence-based medical reference for the prevention of HT.

## METHODOLOGY

### Bibliographic search

The present systematic review and meta-analysis were performed under PRISMA guidelines<sup>[14]</sup>. An online bibliographic search was performed in Pubmed (for English), and China National Knowledge Infrastructure (CNKI) and Wanfang Databases (both for Chinese) by two investigators with key words of 'Vitamin D' and 'Hashimoto thyroiditis'. Studies were considered if they were in English or Chinese, and updated to 27 September 2021. If the abstract displayed investigation about the relationship between Vitamin D and HT, the full text was read in detail by at least two authors (Fig. 1).

### Inclusion criteria

Studies were finally included if they fulfilled the following criteria: (1) included the comparison of HT patient group with

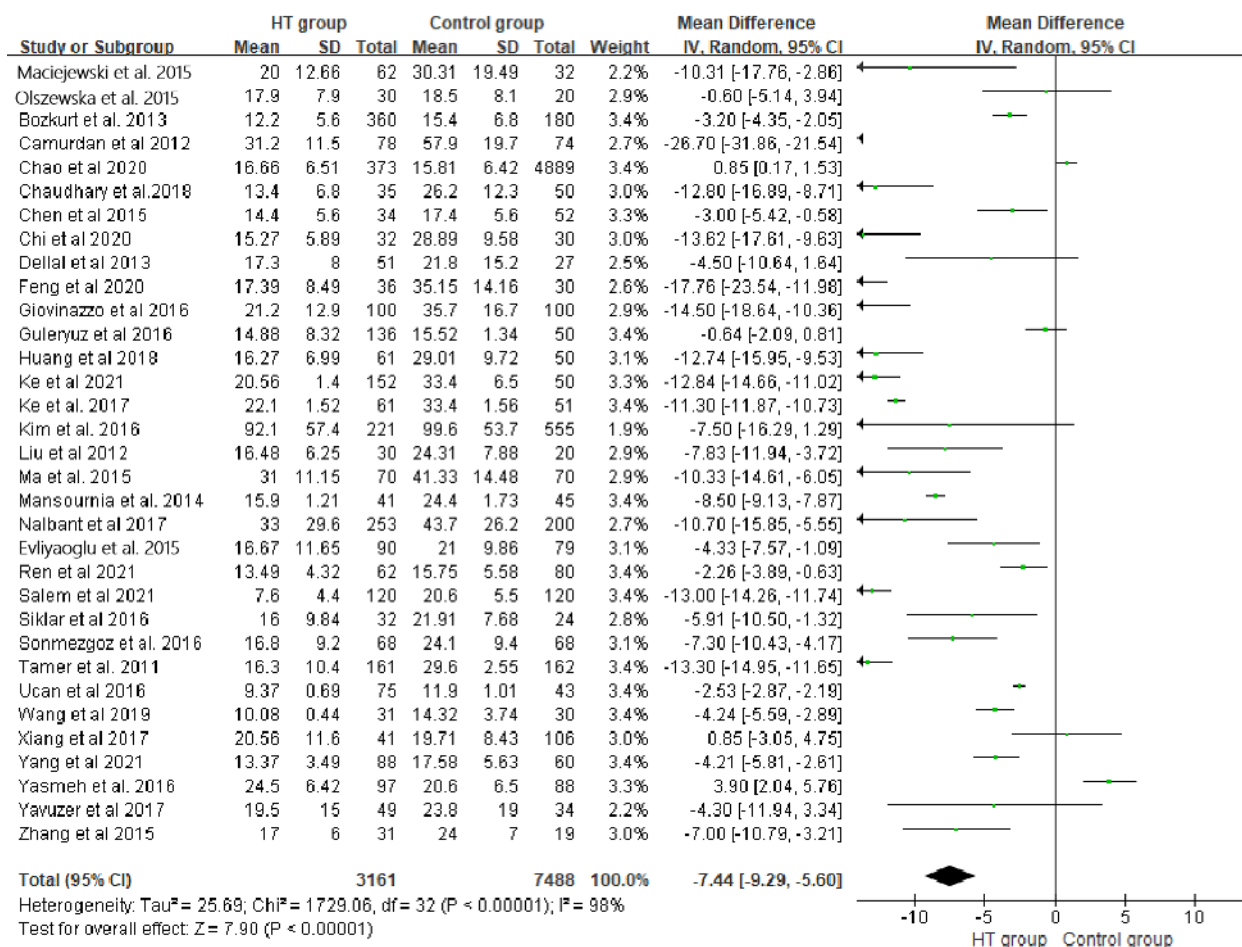


Fig. 1 Forest plot of 25(OH)D<sub>3</sub> level (random model).

a healthy control group; (2) serum levels of 25(OH)D<sub>3</sub> level and/or the quantity of patients with 25(OH)D<sub>3</sub> insufficiency were reported; (3) written in English or Chinese; (4) with a quality score above or equal to 6 according to the coding manual for case-control studies<sup>[15]</sup> assessed by two authors respectively.

**Data extraction**

The following information was extracted from each study by two investigators independently: (1) the first author; (2) year of publication; (3) region; (4) sample size; (5) serum levels of 25(OH)D<sub>3</sub>; (6) cut-off of serum 25(OH)D<sub>3</sub> insufficiency; (7) the quantity of patients with 25(OH)D<sub>3</sub> insufficiency; and (8) quality score. After that, the extracted information was summarized and checked by another two authors.

**Statistical analysis**

RevMan 5.3 (the Cochrane Collaboration) was used to perform a meta-analysis on the data obtained. Firstly, Weighted mean differences (WMD) for continuous variable and Odds Ratios (OR) for binary variables were calculated. Subsequently, statistical heterogeneity was assessed with I<sup>2</sup> test, and the main source of heterogeneity was revealed by subgroup analysis. Lastly, publication bias was evaluated by funnel plots. For statistical analysis above, 'P < 0.05' was considered a significant difference between groups.

**RESULTS**

**Search results**

Online search obtained 336 studies, in which 280 were excluded by abstract screening. Then, of the 56 remaining, 10 were excluded by full text in-detail evaluation; finally 46 studies with 15,336 individuals in total (6,138 HT patients and 9,198 healthy controls) were included into the present study for the systematic review<sup>[16-61]</sup>. Characteristics of included studies are summarized in Table 1.

**Evaluation of indexes**

*Serum 25(OH)D<sub>3</sub> level*

Meta-analysis included 33 studys with 3,161 patients in HT group and 7,488 healthy individuals in the control group for comparison. Random model indicated 25(OH)D<sub>3</sub> levels of HT group were significantly lower than the control group (WMD: -7.44, 95%CI [-9.29, -5.60], P < 0.01). I<sup>2</sup> test (98%) suggested significant heterogeneity in the meta-analysis (Fig. 1). The subgroup meta-analysis basing on 25(OH)D<sub>3</sub> assays in a fixed model revealed similar results (WMDs: -0.55; 95%CI [-0.60, -0.49], P < 0.01), and its significant heterogeneity among subgroups represented by I<sup>2</sup> = 98.3% suggested the difference of 25(OH)D<sub>3</sub> assays was the main source of heterogeneity (Fig. 2.). Lastly, we separated the Chinese studies with 6,639 individuals (HT: 1,102 vs C: 5537) to

**Table 1.** Characteristics of included studies.

Study (published year)	Region	Sample size (HT:C)	25(OH)D <sub>3</sub> Assay method	Serum 25(OH)D <sub>3</sub> level (HT vs C) (ng/mL)	Serum 25(OH)D <sub>3</sub> insufficiency cut off (ng/mL)	Number of 25(OH)D <sub>3</sub> insufficiency (HT:C)	Quality score
Maciejewski et al. 2015 <sup>[23]</sup>	Poland	62/32	ELISA	8.00 ± 5.06 vs 12.12 ± 7.80	<30	61/27	7
Ucan et al. 2016 <sup>[27]</sup>	Turkey	75/43	RIA	9.37 ± 0.69 vs 11.9 ± 1.01	<20	75/36	9
Bozkurt et al. 2013 <sup>[12-17]</sup>	Turkey	360/180	CLS	12.2 ± 5.6 vs 15.4 ± 6.8	<10	150/37	8
Kim 2016 <sup>[20]</sup>	Korea	221/555	CLS	36.84 ± 22.96 vs 39.84 ± 21.48	<30	108/206	8
Sonmezgoz et al. 2016 <sup>[25]</sup>	Turkey	68/68	CLS	16.8 ± 9.2 vs 24.1 ± 9.4	<30	61/54	8
De Pergola et al. 2018 <sup>[18]</sup>	Italy	45/216	CLS	–	<20	31/113	8
Botelho et al. 2018 <sup>[16]</sup>	Brazil	88/71	CLS	26.4 (7.6–48.2) vs 28.6 (13–51.2)	<30	61/39	7
Ma et al. 2015 <sup>[22]</sup>	China	70/70	ELISA	12.40 ± 4.46 vs 16.53 ± 5.79	<30	70/67	7
Yasmeh et al. 2016 <sup>[29]</sup>	America	97/88	CLS	24.5 ± 6.42 vs 20.6 ± 6.5	<30	66/74	7
Xu et al. 2018 <sup>[28]</sup>	China	194/200	CPBA	16.16 (13.72–18.76) vs 23.32 (20.84–25.92)	–	–	7
Kivity et al. 2011 <sup>[21]</sup>	Israel	28/98	CLS	–	<10	22/30	8
Mansournia et al. 2014 <sup>[24]</sup>	Iran	41/45	SC	15.9 ± 1.21 vs 24.4 ± 1.73	<20	34/24	8
Tamer et al. 2011 <sup>[26]</sup>	Turkey	161/162	RIA	16.3 ± 10.4 vs 29.6 ± 2.55	<30	148/102	8
Chaudhary et al. 2018 <sup>[32]</sup>	India	35/50	HPLC	13.39 ± 6.8 vs 26.16 ± 12.28	<20	31/38	8
Evllyaoğlu et al. 2015 <sup>[31]</sup>	Turkey	90/79	HPLC	16.67 ± 11.65 vs 20.99 ± 9.86	<20	80/69	8
Unal et al. 2014 <sup>[30]</sup>	Turkey	254/124	CLS	17.05 (5.4–80) vs 19.9 (9–122.7)	<20	160/-	7
Ke et al. 2017 <sup>[19]</sup>	China	61/51	EBL	22.10 ± 1.52 vs 33.40 ± 1.56	<20	34/12	7
Camurdan et al. 2012 <sup>[33]</sup>	Turkey	78/74	HPLC	31.2 ± 11.5 vs 57.9 ± 19.7	<20	69/24	7
Dellal et al. 2013 <sup>[34]</sup>	Turkey	51/27	RIA	17.3 ± 8.0 vs 21.8 ± 15.2	–	–	6
Siklar et al. 2016 <sup>[35]</sup>	Turkey	32/24	HPLC	16.02 ± 9.84 vs 21.91 ± 7.68	<20	22/10	7
Nalbant et al. 2017 <sup>[36]</sup>	Turkey	253/200	CLS	33 ± 29.6 vs 43.7 ± 26.2	<20	161/111	8
Giovinazzo et al. 2017 <sup>[37]</sup>	Italy	100/100	HPLC	21.2 ± 12.9 vs 35.7 ± 16.7	<20	70/18	7
Guleryuz et al. 2016 <sup>[38]</sup>	Turkey	136/50	HPLC	14.88 ± 8.23 vs 15.52 ± 1.34	–	–	6
Perga et al. 2018 <sup>[39]</sup>	Italy	55/59	CLS	–	<20	37/42	7
Yavuzer et al. 2017	Turkey	49/34	ELISA	19.5 ± 15 vs 23.8 ± 19	–	–	6
Priya et al. 2016	India	25/27	ELISA	14.3 (12.65–17.90) vs 26.2 (21.00–32.8)	–	–	6
Chao et al. 2020 <sup>[42]</sup>	China	373/4889	RIA	16.66 ± 6.51 vs 15.81 ± 6.42	<20	363/4738	9
Feng et al. 2020 <sup>[44]</sup>	China	36/30	ELISA	17.39 ± 8.49 vs 35.15 ± 14.16	–	–	6
Ahi et al. 2020 <sup>[43]</sup>	Iran	633/200	CLS	13.22 (8.1–24.27) vs 20.4 (11.2–29.6)	–	–	7
Liu and Zhang. 2012 <sup>[46]</sup>	China	30/20	RIA	16.48 ± 6.25 vs 24.31 ± 7.88	–	–	7
Xiang et al. 2017 <sup>[47]</sup>	China	41/106	CLS	19.71 ± 8.43 vs 20.56 ± 11.64	<30	38/90	6
Zhang et al. 2015 <sup>[48]</sup>	China	31/19	HPLC	17 ± 6 vs 24 ± 7	–	–	6
Chen et al. 2015 <sup>[45]</sup>	China	34/52	CLS	14.4 ± 5.6 vs 17.4 ± 5.6	<20	29/37	7
Li et al. 2015 <sup>[49]</sup>	China	50/56	–	21.19 (18.40–25.28) vs 24.06 (18.94–33.90)	<30	44/37	6
Cvek et al. 2021 <sup>[50]</sup>	Croatian	461/176	CLS	19.7 (14.4–25.2) vs 17.3 (13.2–22.7)	<20	127/65	7
Salem et al. 2021 <sup>[51]</sup>	Egypt	120/120	ELISA	7.6 ± 4.4 vs 20.6 ± 5.5	<10	120/112	7
Hana et al. 2021 <sup>[52]</sup>	Egypt	112/48	HPLC	10.1 (8.7–11.7) vs 12.0 (9.3–15.6)	<30	101/40	6
Olszewska et al. 2020 <sup>[53]</sup>	Italy	30/20	–	17.9 ± 7.9 vs 18.5 ± 8.1	–	–	6
Rezaee et al. 2017 <sup>[40]</sup>	Iran	51/45	CLS	–	–	–	6
Ren et al. 2021 <sup>[55]</sup>	China	62/80	–	13.49 ± 4.32 vs 15.75 ± 5.85	<30	60/76	6
Huang et al. 2018 <sup>[56]</sup>	China	61/50	CLS	16.27 ± 6.99 vs 29.01 ± 9.72	<20	–	6
Chi et al. 2020 <sup>[57]</sup>	China	32/30	CLS	15.27 ± 5.98 vs 28.89 ± 9.58	–	–	6
Yang et al. 2021 <sup>[58]</sup>	China	88/60	–	13.37 ± 3.49 vs 17.58 ± 5.63	<20	–	6
Ke et al. 2021 <sup>[59]</sup>	China	152/50	CLS	20.56 ± 1.4 vs 33.4 ± 6.5	<20	90/6	7
Wang et al. 2015 <sup>[64]</sup>	China	31/30	ELISA	10.08 ± 0.44 vs 14.32 ± 3.74	–	–	6
Fu et al. 2021 <sup>[61]</sup>	China	334/300	–	16.84 (11.81, 23.39) vs 16.66 (11.98, 22.13)	<30	214/209	7

H: hashimoto thyroiditis group; C: Healthy control group; ELISA: Enzyme Linked Immunosorbent Assay; RIA: Radioimmunoassay; CLS: Chemiluminescent Immunoassay; CPBA: competitive protein binding assay; SC: Solid Chromatography, HPLC: High Performance Liquid Chromatography, EBL: Euglobulin lysis method, –: Non reported.

perform another particle meta-analysis in random model. Result showed that, in the Chinese population, serum 25(OH)D<sub>3</sub> level of HT patients was significantly lower than that of healthy individuals (WMD: –7.04, 95%CI [–10.37,

–3.71],  $P < 0.01$ ). Meanwhile,  $I^2 = 98.0\%$  also suggested a significant heterogeneity (Fig. 3).

#### Prevalence of 25(OH)D<sub>3</sub> insufficiency

A total of 29 studies comprising 11,795 individuals (HT:

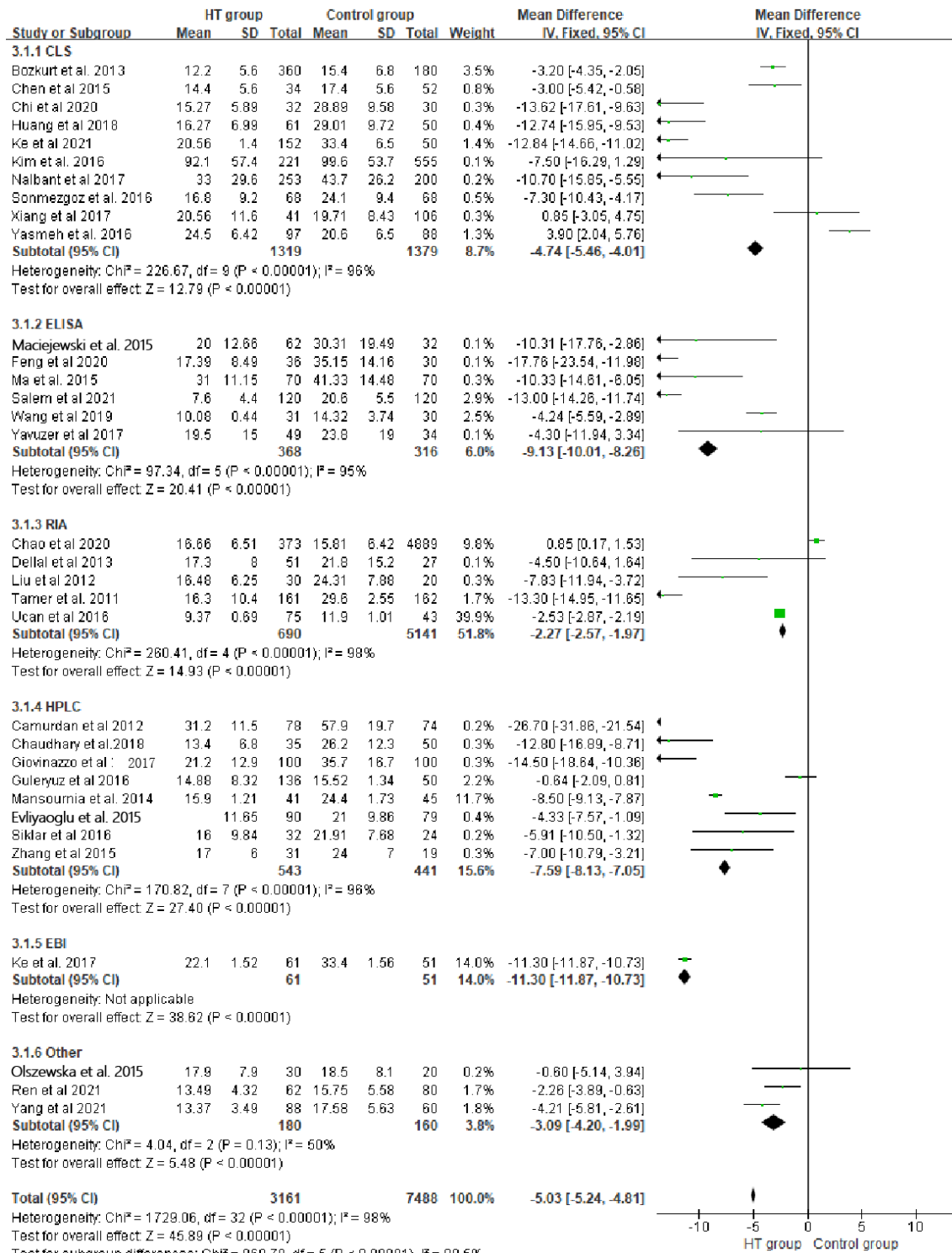


Fig. 2 Subgroup forest plot of 25(OH)D<sub>3</sub> level (fixed model).

3,709 vs C: 8,086) were pooled for OR of 25(OH)D<sub>3</sub> insufficiency. Random model indicated HT patients had higher prevalence of Vitamin D insufficiency compared to healthy individuals (OR: 2.54, 95%CI [1.77, 3.63], P < 0.01). I<sup>2</sup> test (86%) suggested significant heterogeneity in meta-analysis (Fig. 4). Subgroup meta-analysis in a fixed model based on different 25(OH)D<sub>3</sub> insufficiency cut-off also

revealed similar results as above (OR: 1.84; 95%CI [1.64, 2.07], P < 0.01). Meanwhile, I<sup>2</sup> equaled to 93% suggested the main source of heterogeneity was from the different cut-off of 25(OH)D<sub>3</sub> insufficiency (Fig. 5). Chinese studies with 1,373 individuals (HT: 700 vs C: 673) were separated to perform another particle meta-analysis in random model. Results displayed a trend that the HT population had a higher

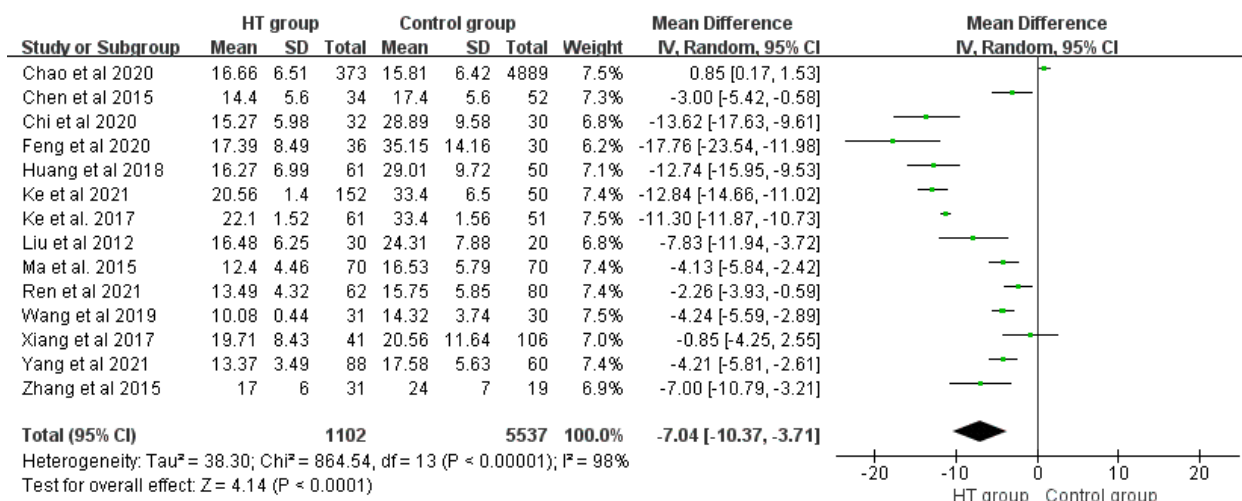


Fig. 3 Forest plot of prevalence of Vitamin D insufficiency (random model).

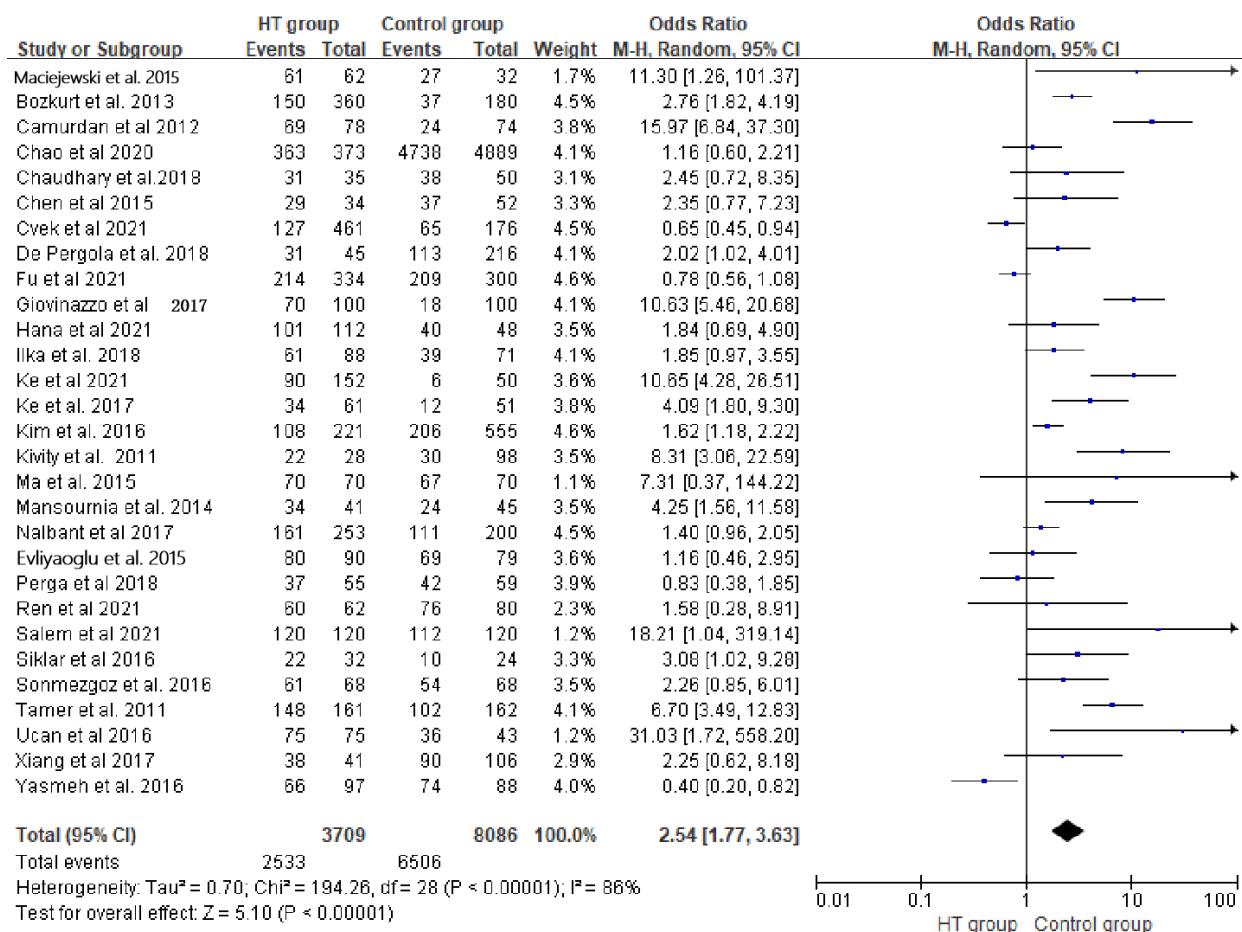


Fig. 4 Subgroup forest plot of prevalence of Vitamin D insufficiency (fixed model).

prevalence of 25(OH)D<sub>3</sub> insufficiency compared to healthy individuals, but it was not statistically significant ( $P > 0.05$ ), and meanwhile significant heterogeneity was indicated by  $I^2$  equal to 91% (Fig. 6).

Publication bias

A funnel plot of serum 25(OH)D<sub>3</sub> level in subgroup analysis

exhibited that the included studies accumulated at the top of the funnel, which suggested that publication bias may exert little adverse effect on the confidence in the meta-analysis (Fig. 7). Similarly, results of the funnel plot suggested low risk of publication bias in prevalence of 25(OH)D<sub>3</sub> insufficiency comparisons (Fig. 8).



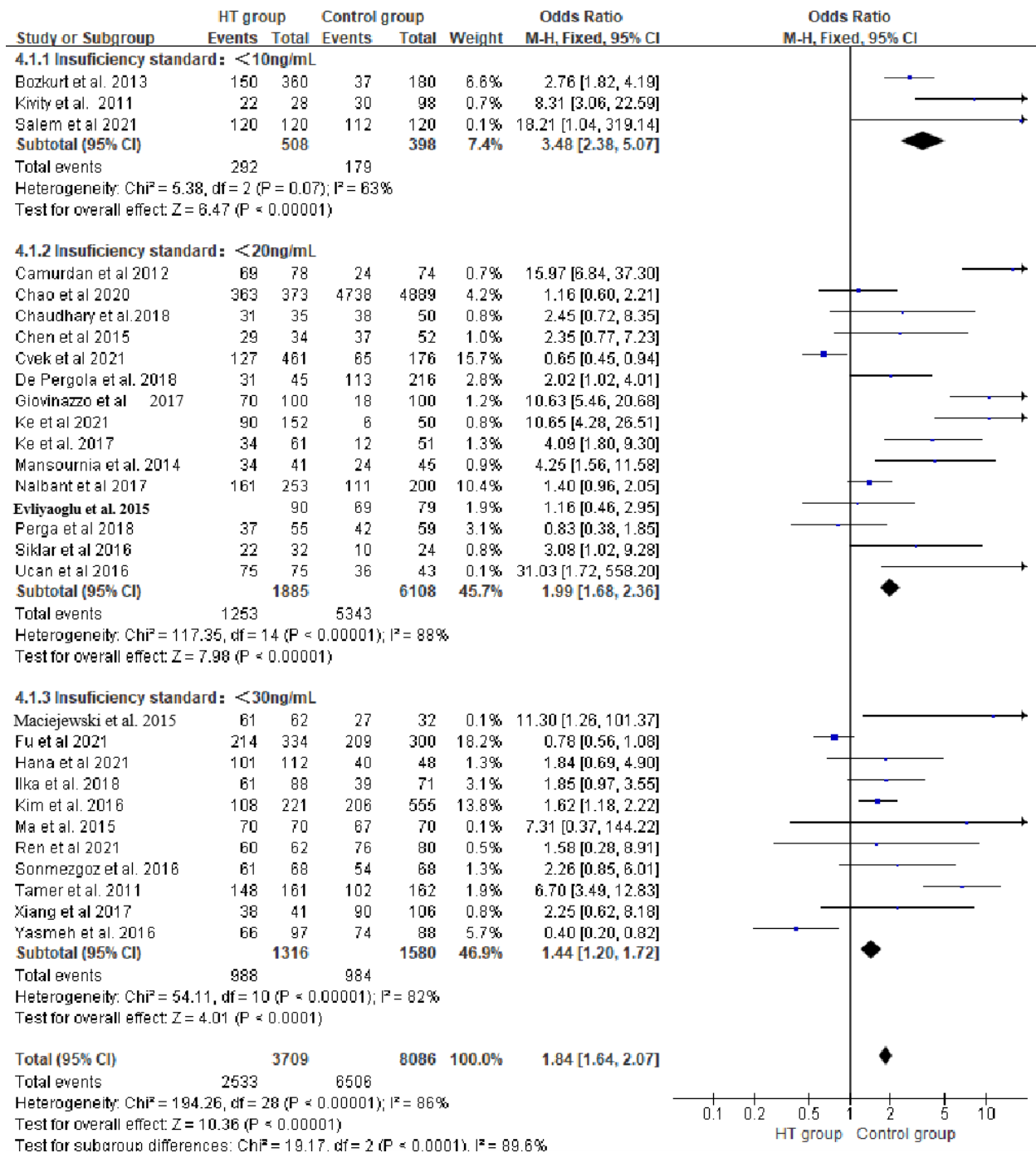


Fig. 5 Forest plot of 25(OH)D<sub>3</sub> level (random model, Chinese studies).

**DISCUSSION**

The present study reinforced the close relationship between Vitamin D insufficiency and HT with methods of systematic review and meta-analysis. To our knowledge, the present systematic review summarized the most related studies to date; among them, Chinese studies, which may be ignored by other foreign researchers, were also included. Hence, we believe our conclusion produce more confident evidence for a relationship between Vitamin D insufficiency and HT.

Although a series of related factors of HT have been revealed, the real etiology has so far not been clearly understood<sup>[62]</sup>. Vitamin D has been proved to closely relate to HT, for it plays a vital role in regulating inflammatory response and maintaining immune balance<sup>[63]</sup>. Multiple epidemiological studies suggested a close relationship between Vitamin D and HT; however, differences in quality, region and population may affect the conclusions. Therefore, high quality systematic review or meta-analysis was still needed to acquire more reliable evidence. Wang et al.<sup>[64]</sup> published a meta-analysis in 2015 to indicate the relationship

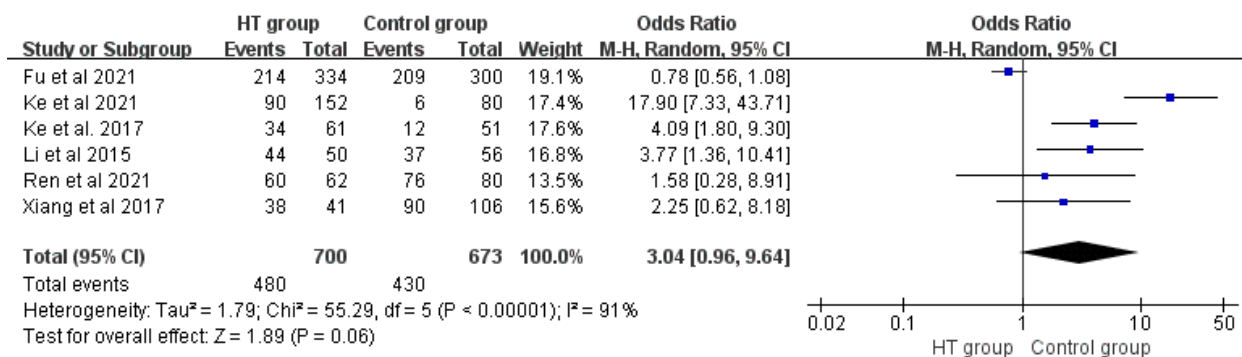


Fig. 6 Forest plot of prevalence of Vitamin D insufficiency (random model, Chinese studies).

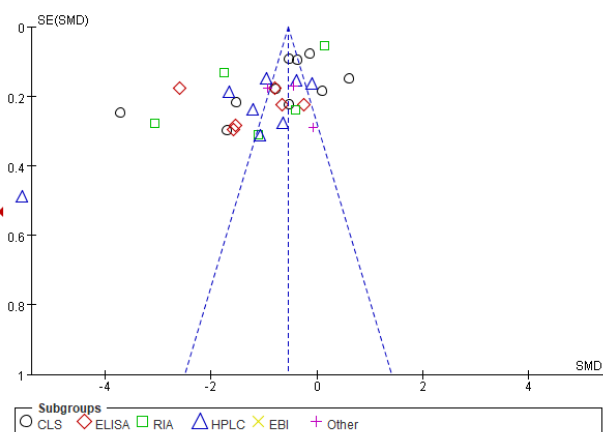


Fig. 7 Funnel plot of 25(OH)D<sub>3</sub> level.

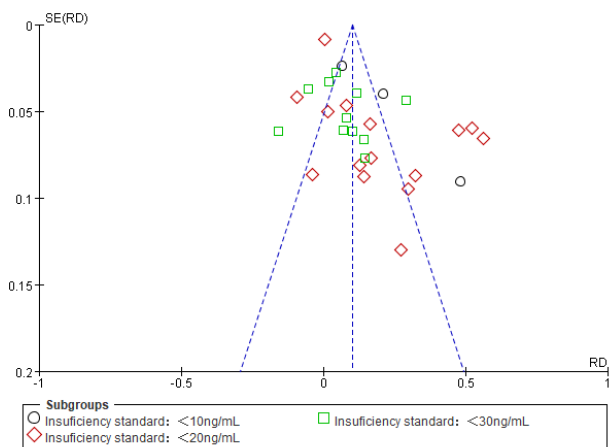


Fig. 8 Funnel plot of quantity of individuals with 25(OH)D<sub>3</sub> insufficiency.

between Vitamin D insufficiency and HT. Štefanić et al. [65] subsequently included more recent studies for meta-analysis and drew a similar conclusion, but their results seemed to weaken the relationship of Vitamin D insufficiency and HT compared with Wang et al.[64]. However, these two studies did not include enough recent Chinese studies which should not be ignored. This may not only decrease the confidence of the conclusions, but also weaken the reliability for Chinese researchers. To fulfill this deficiency, we included Chinese studies into the present systematic review and meta-analysis.

As anticipated, the general results including serum 25(OH)D<sub>3</sub> level and quantity of individuals with Vitamin D insufficiency displayed similar results to Wang et al.[64] and Štefanić et al.[65], which meant the relationship of Vitamin D insufficiency and HT also existed in the Chinese population. We next separated the Chinese studies to perform a particle meta-analysis. The particle result of serum 25(OH)D<sub>3</sub> levels of Chinese HT patients were significantly lower than healthy individuals generally, but its difference was less (-7.05 vs -7.44). However, concerning the prevalence of Vitamin D insufficiency, Chinese HT patients were not likely to have more Vitamin D insufficiency cases compared to healthy individuals, suggesting the relationship between Vitamin D insufficiency and HT in the Chinese population may not be as strong as in the global population. Note that, the Chinese population in the present study was only a small part of the total, the negative result maybe due to the small sample size. Further studies with larger sample sizes and high quality investigating the prevalence of Vitamin D insufficiency in the HT population are necessary in China in the future.

With regards to the studies included in the systematic review, most studies reported lower serum 25(OH)D<sub>3</sub> levels and higher prevalence of Vitamin D insufficiency in HT patients compared to healthy individuals. The present study had drawn a similar conclusion, but the results contained significant heterogeneity. According to the systematic review, this heterogeneity may be due to the difference of 25(OH)D<sub>3</sub> assays, thus we performed an analysis which separated studies with the same assay into several sub-groups. Results similarly indicated lower serum 25(OH)D<sub>3</sub> levels in HT patients, and the most significant heterogeneity among sub-groups (I<sup>2</sup> = 99.5%), which hinted that the heterogeneity was mainly caused by the difference in 25(OH)D<sub>3</sub> assays. In parallel with lower serum 25(OH)D<sub>3</sub> levels, HT patients were at higher risk of 25(OH)D<sub>3</sub> insufficiency, indicated by our meta-analysis. Meanwhile, significant heterogeneity was indicated owing to the difference of serum 25(OH)D<sub>3</sub> insufficiency criteria. With regard to the publication bias, we determined that heterogeneity would bring significant publication bias displayed by the funnel plot, but the results showed that the included studies accumulating at the top of the funnel; this suggested the publication bias exerted little influence on our results. However, our study also had limitations, as 25(OH)D<sub>3</sub> level in a population could be affected by many other factors such as sunshine duration, season, area, economy, and

education, which was not considered in our study. These factors may affect the conclusions of epidemiological research, and bring bias to the meta-analysis. Therefore, we can only indicate that vitamin D insufficiency was related to HT. Whether vitamin D insufficiency could lead to HT should be further investigated by biological research in the future.

Until recently, further investigations focussed on the HT mechanism in which vitamin D was involved. As is well understood, T lymphocytes including Th1, Th2 and Th17 cells infiltrate the thyroid gland due to immunological disorders in HT patients. Vitamin D can inhibit the differentiation of Th1 cells, and the production of inflammatory cytokines such as TNF- $\alpha$ , INF- $\gamma$ . It could also suppress inflammatory Th1, but induce anti-inflammatory Th2 which produced anti-inflammatory cytokines such as IL-4 and IL-5<sup>[66]</sup>. Furthermore, the Th17 cells with their production of IL-17A could also be inhibited by Vitamin D at the transcriptional level<sup>[67]</sup>. On the other hand, vitamin D could increase the proportion of Treg cells to exert immune regulation<sup>[68]</sup>. Taken together, vitamin D may have the potential to prevent HT. However, clinical studies have shown contradictory results: Chahardoli et al.<sup>[69]</sup> reported activated vitamin D supplementation can decrease TSH and TG-Ab antibodies levels in HT patients, but another study showed that activated vitamin D supplementation had no effect on improving HT<sup>[70]</sup>. To our knowledge, the studies mentioned above may ignore the vitamin D receptors polymorphism. Vitamin D receptors in the thyroid gland have single nucleotide polymorphism and most typical Apal, Bsml, Fokl and Taql single nucleotide variations have been shown to be closely related to autoimmune diseases<sup>[71]</sup>. Therefore, more prospective studies are needed to confirm the preventive effect of vitamin D on HT.

## CONCLUSIONS

In conclusion, the present systematic review and meta-analysis strengthened the relationship between vitamin D insufficiency and HT. HT patients potentially had higher propensity for having lower serum 25(OH)D<sub>3</sub> levels compared to healthy individuals. Clinical staff may have to carefully consider the possibility of vitamin D insufficiency in HT patients.

## Conflict of interest

The authors declare that they have no conflict of interest.

## Dates

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